

Branching processes in random environment.

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Branching diffusions in a random environment

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Recall that we define a sequence of **GWRE's** as follows: Fix $n \geq 1$ and let $\mathbf{e}^{(n)} = (\mathbf{Q}_i^{(n)})_{i \geq 1}$ a sequences of i.i.d. r.v.'s. Conditioned on $\mathbf{e}^{(n)}$ the BPRE $(Z_i^{(n)})_{i \geq 0}$ is defined as

$$Z_{i+1}^{(n)} = \sum_{j=1}^{Z_i^{(n)}} X_{j,i}, \quad i \geq 0,$$

where $Z_0^{(n)}$ is independent of $\mathbf{e}^{(n)}$ and where $(X_{j,i}^{(n)})_{j,i \geq 0}$ conditioned on $\mathbf{e}^{(n)}$ are i.i.d. with common distribution

$$\mathbb{P}\left(X_{j,i}^{(n)} = k \mid \mathbf{e}^{(n)}\right) = \mathbf{Q}_i^{(n)}(k), \quad j, i, k \geq 0.$$

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$$\mathbf{m}_i^{(n)} = \sum_{k \geq 0} k \mathbf{Q}_i^{(n)}(k).$$

Let $Z_t^{(n)} = Z_{[t]}^{(n)}$ be a continuous time version of the GWRE and $(S_t^{(n)})_{t \geq 0}$ be its associated random walk which is defined by

$$S_t^{(n)} = \sqrt{n} \sum_{i=0}^{[t]-1} \log \left(\mathbf{m}_i^{(n)} \right), \quad t \geq 0.$$

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If $Z_0^{(n)}/n \rightarrow Z_0$ in law. Hence under some *technical* assumptions we have

$$\left(\frac{Z_{tn}^{(n)}}{n}, \frac{S_{tn}^{(n)}}{\sqrt{n}} \right)_{t \geq 0} \rightarrow (Z_t, S_t)_{t \geq 0}$$

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in the sense of Skorokhod. Moreover

$$Z_t = Z_0 + a \int_0^t Z_s ds + \int_0^t \sqrt{2\gamma^2 Z_s} dB_s + \int_0^t Z_s dS_s,$$

where $S_t = \beta t + \sigma W_t$, $t \geq 0$, $\beta \in \mathbb{R}$ and B and W are two independent standard Brownian motions.

We are interested on the strong solutions of

$$Z_t = Z_0 + (a + \beta) \int_0^t Z_s ds + \int_0^t \sqrt{2\gamma^2 Z_s} dB_s + \sigma \int_0^t Z_s dW_s,$$

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For the moment, we assume that there is a unique strong solution of the previous SDE and for simplicity, we study the case where $\sigma = \alpha = \beta = 0$. For convenience, we denote by Y the unique strong solution of

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Since the previous SDE has a unique strong solution of the above equation, then Y is Markovian and satisfies the **branching property**.

Indeed, let \mathbb{P}_x denotes the law of Y such that $\mathbb{P}_x(Y_0 = x) = 1$. Then we would like to show that \mathbb{P}_{x+y} is equal in law to the convolution of \mathbb{P}_x and \mathbb{P}_y .

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Let β be a third BM independent of B and \tilde{B} . The process

$$C_t = \int_0^t \mathbf{1}_{\{X_s > 0\}} \frac{\sqrt{2\gamma^2 Y_s} dB_s + \sqrt{2\gamma^2 \tilde{Y}_s} d\tilde{B}_s}{\sqrt{X_s}} + \int_0^t \mathbf{1}_{\{X_s = 0\}} d\beta_s,$$

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is a linear BM since $\langle C \rangle_t = t$, implying

$$X_t = x + y + \int_0^t 2\sqrt{\gamma^2 X_s} dC_s.$$

Take $\lambda \geq 0$. From Itô's formula, one can deduce that

$$M_s = \exp \left\{ \frac{-\lambda Y_s}{1 + \gamma^2 \lambda (t - s)} \right\}, \quad \text{for } 0 \leq s \leq t,$$

is a martingale.

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$$\mathbb{E}_x[e^{-\lambda Y_t}] = \exp \left\{ \frac{-\lambda x}{1 + \gamma^2 \lambda t} \right\}.$$

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Therefore

$$\mathbb{E}_x[e^{-\lambda Y_t}] = \exp \left\{ \frac{-\lambda x}{1 + \gamma^2 \lambda t} \right\}.$$

Hence, the survival probability satisfies

$$\mathbb{P}_x(Y_t > 0) = 1 - e^{-x(\gamma^2 t)^{-1}}.$$

Recall that

$$Z_t = Z_0 + (a + \beta) \int_0^t Z_s ds + \int_0^t \sqrt{2\gamma^2 Z_s} dB_s + \sigma \int_0^t Z_s dW_s. \quad (1)$$

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Proposition

Let $\mathbf{m} = \beta + a - \sigma^2/2$ and define

$$\tau(t) = \int_0^t e^{-\sigma W_s - \mathbf{m}s} ds.$$

Hence the process $(e^{\sigma W_t + \mathbf{m}t} Y_{\tau(t)}, t \geq 0)$ is a weak solution of (1).

Proof:

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$$\langle e^{\sigma W \cdot + \mathbf{m} \cdot}, Y_{\tau(\cdot)} \rangle_t = 0.$$

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Applying again Itô's formula and from the definition of Y as an SDE, we deduce

$$Z_t = Y_0 + \sigma \int_0^t Z_s dW_s + (\mathbf{m} + \sigma^2/2) \int_0^t Z_s ds + \int_0^t e^{\sigma W_s + \mathbf{m}s} \sqrt{2\gamma^2 Y_{\tau(s)}} dB_{\tau(s)}.$$

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The proof follows by noting that $\mathbf{m} + \sigma^2/2 = \beta + a$ and

$$M_t = \int_0^t \sqrt{e^{\sigma W_s + \mathbf{m}s}} dB_{\tau(s)}$$

is a Brownian motion. The latter holds since $\langle M \rangle_t = t$.

Corollary

Let Z be a branching diffusion in a Brownian random environment (BDRE). Then a.s.

$$\mathbb{E}_x[e^{-\lambda Z_t} | B] = \exp \left\{ -x \left(\lambda^{-1} + \gamma^2 \int_0^t e^{-\sigma W_s - \mathbf{m}s} ds \right)^{-1} \right\}.$$

In particular

$$\mathbb{P}(Z_t > 0 | B) = 1 - \exp \left\{ -x \left(\gamma^2 \int_0^t e^{-\sigma W_s - \mathbf{m}s} ds \right)^{-1} \right\}.$$

Hence the probability of extinction depends on the law of

$$A_t^{(\eta)} := \int_0^t \exp \left\{ 2(\eta s + W_s) \right\} ds, \quad t \in [0, \infty).$$

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According to Matsumoto and Yor (2003), for all $t, u \in (0, \infty)$ and $x \in \mathbb{R}$, we have

$$\mathbb{P}(A_t^{(\eta)} \in du | W_t + \eta t = x) = \frac{\sqrt{2\pi t}}{u} e^{\frac{x^2}{2t}} e^{-\frac{1}{2u} (1+e^{2x})} \theta_{e^{x/u}}(t) du,$$

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According to Bertoin and Yor (05), $A_\infty^{(\eta)} = \lim_{t \rightarrow \infty} A_t^{(\eta)}$ is a finite a.s. whenever $\eta < 0$. Moreover ,

$$A_\infty^{(\eta)} \quad \text{has the same law as} \quad \left(2\Gamma_{-\eta}\right)^{-1},$$

where Γ_v is Gamma r.v. with parameter v .

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Theorem

Let $(Z_t, t \geq 0)$ be a BDRE. Then for all $z > 0$,

i) (Supercritical case) If $\mathbf{m} > 0$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}_z(Z_t > 0) = 1 - \sum_{n=0}^{\infty} \frac{(-z)^n}{n!} \mathbf{k}^n \frac{\Gamma(n - \eta)}{\Gamma(-\eta)}.$$

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ii) (Critical case) If $\mathbf{m} = 0$, we have

$$\lim_{t \rightarrow \infty} \sqrt{t} \mathbb{P}_z(Z_t > 0) = -\frac{\sqrt{2}}{\sqrt{\pi\sigma}} \sum_{n=1}^{\infty} \frac{(-z)^n}{n!} \mathbf{k}^n \Gamma(n).$$

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iii) (Weakly subcritical) If $\mathbf{m} \in (-\sigma^2, 0)$, we have

$$\lim_{t \rightarrow \infty} t^{\frac{3}{2}} e^{\frac{\mathbf{m}^2 t}{2\sigma^2}} \mathbb{P}_z(Z_t > 0) = \frac{8}{\sigma^3} \int_0^\infty (1 - e^{-\mathbf{k}zv}) \phi_\eta(v) dv,$$

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iv) (Intermediately subcritical) If $\mathbf{m} = -\sigma^2$, we have

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v) (Strongly subcritical) If $\mathbf{m} < -\sigma^2$, we have

$$\lim_{t \rightarrow \infty} e^{\frac{\sigma^2 t}{2}} e^{-(\mathbf{m} + \sigma^2)t} \mathbb{P}_z(Z_t > 0) = z \mathbf{k}(\eta - 2).$$

Branching diffusions with catastrophes

We are now interested in the following SDE

$$Z_t = Z_0 + a \int_0^t Z_s ds + \int_0^t \sqrt{2\gamma^2 Z_s} dB_s + \int_0^t Z_{s-} dS_s,$$

where $a \in \mathbb{R}$ and S is a Compound Poisson process independent of B whose jump distribution is supported on $(-1, 0)$ and jumps with rate $r > 0$.

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Recall that the scaling limit of GW processes is a Feller diffusion with growth a and diffusion part γ^2 . The latter is also the scaling limit of birth and death processes and it gives a natural model for populations which die and multiply fast, randomly, without interaction.

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The process Z was considered by Bansaye and Tran (2011) to model parasites growing in dividing cells.

The cell divides at constant rate r and a random fraction $\Theta \in (0, 1)$ of parasites enters the first daughter cell, whereas the remainder enters the second daughter cell. Following the infection in a cell line, the parasites grow as a Feller diffusion process and undergo a catastrophe when the cell divides.

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We denote by N_t and N_t^* the numbers of cells and infected cells at time t , respectively. We say that the cell population recovers when the asymptotic proportion of contaminated cells vanishes.

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Observe that if there is one infected cell at time 0, $\mathbb{E}[N_t] = e^{rt}$ and $\mathbb{E}[N_t^*] = e^{rt} \mathbb{P}(Z_t > 0)$, since N is a Yule process.

It is important to note that the process Z can be written as follows

$$Z_t = Z_0 + a \int_0^t Z_s ds + \int_0^t \sqrt{2\gamma^2 Z_s} dB_s + \int_0^t \int_{(0,1)} (\theta - 1) Z_{s-} M(ds, d\theta), \quad (2)$$

where $M(ds, d\theta)$ a Poisson random measure with intensity $rds \mathbb{P}(\Theta \in d\theta)$.

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Proposition

Let

$$K_t = at + \int_0^t \int_{(0,1)} \log \theta M(ds, d\theta),$$

and define

$$\tau(t) = \int_0^t e^{-K_s} ds.$$

Hence the process $(e^{K_t} Y_{\tau(t)}, t \geq 0)$ is a weak solution of (2).

Proof:

Proof: Again, we denote $Z_t = e^{Kt} Y_{\tau(t)}$. Applying Itô's formula, we deduce

$$Z_t = Y_0 + \int_0^t Y_{\tau(s)} de^{\sigma Ks} + \int_0^t e^{\sigma Ks} dY_{\tau(s)}.$$

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From the independence of B and M , we have $[e^{K\cdot}, Y_{\tau(\cdot)}]_t = 0$.

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From the independence of B and M , we have $[e^{K\cdot}, Y_{\tau(\cdot)}]_t = 0$.

By noting that

$$e^{Kt} = 1 + a \int_0^t e^{Ks} ds + \int_0^t \int_{(0,1)} e^{Ks - (\theta - 1)M(ds, d\theta)}.$$

and Itô's formula, we deduce

$$Z_t = Y_0 + a \int_0^t Z_s ds + \int_0^t \int_{(0,1)} (\theta - 1) Z_{s-} M(ds, d\theta) + \int_0^t e^{Ks} \sqrt{2\gamma^2 Y_{\tau(s)}} dB_{\tau(s)}.$$

Proof: Again, we denote $Z_t = e^{Kt} Y_{\tau(t)}$. Applying Itô's formula, we deduce

$$Z_t = Y_0 + \int_0^t Y_{\tau(s)} de^{\sigma K_s} + \int_0^t e^{\sigma K_s} dY_{\tau(s)}.$$

From the independence of B and M , we have $[e^{K\cdot}, Y_{\tau(\cdot)}]_t = 0$.

By noting that

$$e^{Kt} = 1 + a \int_0^t e^{Ks} ds + \int_0^t \int_{(0,1)} e^{Ks - (\theta - 1)M(ds, d\theta)}.$$

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The proof follows by observing that

$$M_t = \int_0^t \sqrt{e^{Ks}} dB_{\tau(s)}$$

is a Brownian motion since $\langle M \rangle_t = t$.

Corollary

Let Z be a branching diffusion with catastrophes. Then a.s.

$$\mathbb{E}_x[e^{-\lambda Z_t} | K] = \exp \left\{ -x \left(\lambda^{-1} + \gamma^2 \int_0^t e^{-K_s} ds \right)^{-1} \right\}.$$

In particular

$$\mathbb{P}(Z_t > 0 | K) = 1 - \exp \left\{ -x \left(\gamma^2 \int_0^t e^{-K_s} ds \right)^{-1} \right\}.$$

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Unfortunately there is not an explicit form for the density of the exponential functional of K . The only case known who has explicit density is the case of Brownian motion with drift.

Assume that the Laplace exponent of the process K is well-defined for some positive real number, i.e.

$$\psi(\lambda) = \log \mathbb{E}[e^{\lambda K_1}] < \infty \quad \text{for } \lambda \in [0, \theta^+),$$

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which converges in distribution to a positive finite limit proportional to x_0 . Moreover,

$$\frac{1}{t} \log \mathbb{P}_x(Z_t > 0 \mid K) \rightarrow \psi'(0),$$

in probability.

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(iii) If $\psi'(1) > 0$ (*weakly s.r.*) and $\theta^+ > 2$, then there exists $c_3 > 0$ such that

$$\mathbb{P}_x(Z_t > 0) \sim c_3 t^{-3/2} e^{t\psi(\tau)+g}, \quad \text{as } t \rightarrow \infty,$$

where τ is s.t. $\psi(\tau) = \min_{0 < s < 1} \psi(s)$.

Proposition

b) If $\psi'(0) = 0$ (critical case) and $\theta^+ > 1$, then for every $x > 0$, there exists $c_4 > 0$ such that

$$\mathbb{P}_x(Z_t > 0) \sim c_4 t^{-1/2}, \quad \text{as } t \rightarrow \infty.$$

CB-processes.

A continuous-state branching process (or CB-process) is a non-negative valued strong Markov process with probabilities $(\mathbb{P}_x, x \geq 0)$ such that for any $x, y \geq 0$, \mathbb{P}_{x+y} is equal to the convolution of \mathbb{P}_x and \mathbb{P}_y .

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In particular,

$$\mathbb{E}_x \left[e^{-\lambda Y_t} \right] = \exp\{-xu_t(\lambda)\}, \quad \text{for } \lambda \geq 0,$$

for some function $u_t(\lambda)$.

The function $u_t(\lambda)$ is determined by the integral equation

$$\int_{u_t(\lambda)}^{\lambda} \frac{1}{\psi(u)} du = t$$

where ψ (**branching mechanism** of Y) satisfies the Lévy-Khinchine formula

$$\psi(\lambda) = -a\lambda + \gamma^2 \lambda^2 + \int_{(0, \infty)} (e^{-\lambda x} - 1 + \lambda x \mathbf{1}_{\{x < 1\}}) \mu(dx),$$

where $a \in \mathbb{R}$, $\gamma \geq 0$ and μ is a σ -finite measure such that

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Observe $\mathbb{E}_x[Y_t] = xe^{-\psi'(0^+)t}$. Hence, in respective order, a CB-process is called **supercritical**, **critical** or **subcritical** accordingly as $\psi'(0^+) < 0$, $\psi'(0^+) = 0$ or $\psi'(0^+) > 0$.

The probability of extinction is given by

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A CB-process Y with branching mechanism ψ has a finite time **extinction** almost surely if and only if

$$\int^{\infty} \frac{du}{\psi(u)} < \infty \quad \text{and} \quad \psi'(0+) \geq 0.$$

A CB-process can also be defined as the unique non-negative strong solution of the stochastic differential equation

$$\begin{aligned}
 Y_t = & Y_0 + a \int_0^t Y_s ds + \int_0^t \sqrt{2\gamma^2 Y_s} dB_s \\
 & + \int_0^t \int_{(0,1)} \int_0^{Y_{s-}} z \tilde{N}(ds, dz, du) + \int_0^t \int_{[1,\infty)} \int_0^{Y_{s-}} z N(ds, dz, du),
 \end{aligned}$$

where $B = (B_t, t \geq 0)$ is a standard Brownian motion, N is a Poisson random measure independent of B , with intensity $ds \otimes \mu(dz) \otimes du$ and \tilde{N} is its compensated version.

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